

Multilevel, multivariate, and network meta-analysis with the *metafor* package in R

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Purpose of Talk

- describe how multilevel and multivariate structures can arise in meta-analytic data
- illustrate how to fit multilevel, multivariate, and network meta-analyses with the *metafor* package in R

Meta-Analytic Data

- $i = 1, \dots, k$ studies
- have y_i and corresponding v_i
- assume:

$$y_i | \theta_i \sim N(\theta_i, v_i)$$

- and independence of the estimates (for now)
- approx. 95% CI for θ_i : $y_i \pm 1.96\sqrt{v_i}$

Example: BCG Vaccine

- effectiveness of the Bacillus Calmette-Guérin (BCG) vaccine against tuberculosis (TB)
- for each study, can compare the proportion of TB positive cases in the vaccinated versus the non-vaccinated group

Example: BCG Vaccine

| | | Tuberculosis | | |
|------------|----|--------------|----------|--|
| | | Positive | Negative | |
| Vaccinated | 4 | 119 | 123 | |
| | 11 | 128 | 139 | |

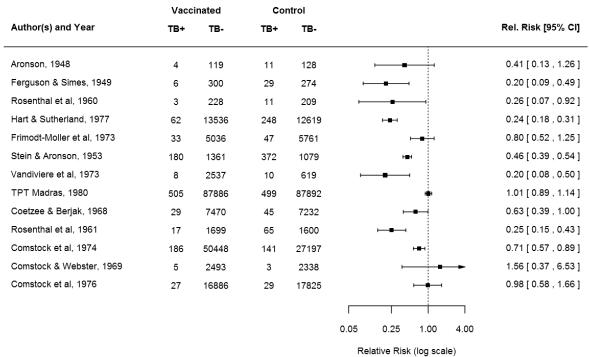
$p_T = 4/123 = .0325$
 $p_C = 11/139 = .0791$
 $RR = \frac{4/123}{11/139} = .41$

$y = \ln[RR] = \ln\left[\frac{4/123}{11/139}\right] = -.89$
 $v = \frac{1}{4} - \frac{1}{123} + \frac{1}{11} - \frac{1}{139} = .326$

Example: BCG Vaccine

| Study | Year | RR | $y = \ln(RR)$ | v | Allocation | Latitude |
|-------|------|------|---------------|------|------------|----------|
| 1 | 1948 | 0.41 | -0.89 | .326 | random | 44 |
| 2 | 1949 | 0.20 | -1.59 | .195 | random | 55 |
| 3 | 1960 | 0.26 | -1.35 | .415 | random | 42 |
| 4 | 1977 | 0.24 | -1.44 | .020 | random | 52 |
| 5 | 1973 | 0.80 | -0.22 | .051 | alternate | 13 |
| 6 | 1953 | 0.46 | -0.79 | .007 | alternate | 44 |
| 7 | 1973 | 0.20 | -1.62 | .223 | random | 19 |
| 8 | 1980 | 1.01 | 0.01 | .004 | random | 13 |
| 9 | 1968 | 0.63 | -0.47 | .056 | random | 27 |
| 10 | 1961 | 0.25 | -1.37 | .073 | systematic | 42 |
| 11 | 1974 | 0.71 | -0.34 | .012 | systematic | 18 |
| 12 | 1969 | 1.56 | 0.45 | .533 | systematic | 33 |
| 13 | 1976 | 0.98 | -0.02 | .071 | systematic | 33 |

Example: BCG Vaccine



Standard Random-Effects Model

$$\begin{aligned}
 y_i &= \mu && \text{average true outcome} \\
 &+ u_i && \text{random effect that makes the true outcome} \\
 &+ e_i && \text{for a particular study larger/smaller by some} \\
 & && \text{amount (heterogeneity between studies)} \\
 & && \text{sampling error}
 \end{aligned}$$

$$e_i \sim N(0, v_i) \quad u_i \sim N(0, \tau^2)$$

Marginal Variance-Covariance Matrix

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \\ y_5 \\ \vdots \\ y_k \end{bmatrix} = \begin{bmatrix} \tau^2 + v_1 & & & & & & \\ & \tau^2 + v_2 & & & & & \\ & & \tau^2 + v_3 & & & & \\ & & & \tau^2 + v_4 & & & \\ & & & & \tau^2 + v_5 & & \\ & & & & & \ddots & \\ & & & & & & \tau^2 + v_k \end{bmatrix}$$

Meta-Analysis with R

- **metafor:** meta-analysis package for R
- install with: `install.packages("metafor")`
- load with: `library(metafor)`
- comments start with `#`

```

> ### load BCG vaccine data
> dat <- get(data(dat.bcg))
>
> ### show data
> dat

```

| trial | author | year | treated | | | | control | alloc | |
|-------|--------|----------------------|---------|------|-------|------|---------|-------|------------|
| | | | tpos | tneg | cpos | cneg | | | |
| 1 | 1 | Aronson | 1948 | 4 | 119 | 11 | 128 | 44 | random |
| 2 | 2 | Ferguson & Simes | 1949 | 6 | 300 | 29 | 274 | 55 | random |
| 3 | 3 | Rosenthal et al | 1960 | 3 | 228 | 11 | 209 | 42 | random |
| 4 | 4 | Hart & Sutherland | 1977 | 62 | 13536 | 248 | 12619 | 52 | random |
| 5 | 5 | Frimodt-Møller et al | 1973 | 33 | 5036 | 47 | 5761 | 13 | alternate |
| 6 | 6 | Stein & Aronson | 1953 | 180 | 1361 | 372 | 1079 | 44 | alternate |
| 7 | 7 | Vandiviere et al | 1973 | 8 | 2537 | 10 | 619 | 19 | random |
| 8 | 8 | TPT Madras | 1980 | 505 | 87886 | 499 | 87892 | 13 | random |
| 9 | 9 | Coetze & Berjak | 1968 | 29 | 7470 | 45 | 7232 | 27 | random |
| 10 | 10 | Rosenthal et al | 1961 | 17 | 1699 | 65 | 1600 | 42 | systematic |
| 11 | 11 | Comstock et al | 1974 | 186 | 50448 | 141 | 27197 | 18 | systematic |
| 12 | 12 | Comstock & Webster | 1969 | 5 | 2493 | 3 | 2338 | 33 | systematic |
| 13 | 13 | Comstock et al | 1976 | 27 | 16886 | 29 | 17825 | 33 | systematic |

Computing Observed Outcomes

- can of course use external software for data management and preparations
- to compute outcomes: `escalc()` command
- basic syntax:

```
dat <- escalc(measure="", ..., data=dat)
```

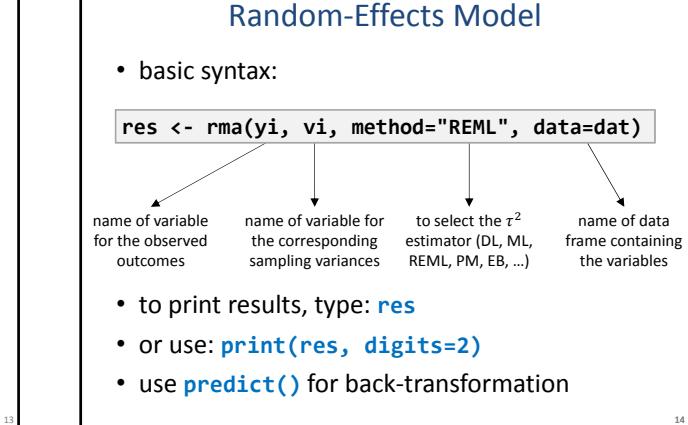
to specify the outcome measure (RD, RR, OR, SMD, ROM, PLO, ...)

to specify the variables needed to compute the observed outcomes

```
> ### calculate log relative risks and sampling variances
> dat <- escalc(measure="RR", ai=tpos, bi=tneg,
+ ci=cpos, di=cneg, data=dat)
> dat

  trial author year ... yi vi
  1      Aronson 1948 ...
  2      Ferguson & Simes 1949 ...
  3      Rosenthal et al 1960 ...
  4      Hart & Sutherland 1977 ...
  5      Frimodt-Møller et al 1973 ...
  6      Stein & Aronson 1953 ...
  7      Vandiviere et al 1973 ...
  8      TPT Madras 1980 ...
  9      Coetzee & Berjak 1968 ...
 10     Rosenthal et al 1961 ...
 11     Comstock et al 1974 ...
 12     Comstock & Webster 1969 ...
 13     Comstock et al 1976 ...

log relative risks and
sampling variances
```



```
> ### fit random-effects model
> res <- rma(yi, vi, data=dat)
> res

Random-Effects Model (k = 13; tau^2 estimator: REML)

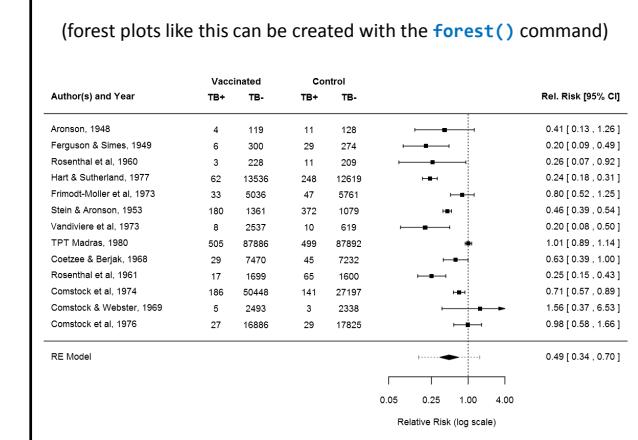
tau^2 (estimated amount of total heterogeneity): 0.3132
tau (square root of estimated tau^2 value):   0.5597
I^2 (total heterogeneity / total variability):  92.22%
H^2 (total variability / sampling variability): 12.86

Test for Heterogeneity:
Q(df = 12) = 152.2330, p-val < .0001

Model Results:

estimate   se   zval   pval   ci.lb   ci.ub
-0.7145  0.1798 -3.9744 < .0001  -1.0669  -0.3622

> ### estimated average relative risk (and 95% CI/CR)
> pred(res, transf=exp, digits=2)
pred ci.lb ci.ub cr.lb cr.ub
 0.49  0.34  0.70  0.15  1.55
→ cr.lb/cr.ub = bounds of a 95% credibility/prediction interval
```



Mixed-Effects Meta-Regression Model

- can include moderators/predictors/covariates in the model (to account for heterogeneity)
- mixed-effects meta-regression model:
 - $y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + u_i + e_i$
 - $u_i \sim N(0, \tau^2)$ (but now 'residual' heterogeneity)
 - $e_i \sim N(0, v_i)$

Mixed-Effects Meta-Regression Model

- basic syntax as before, but now:
- ```
res <- rma(yi, vi, mods = ~ var1, data=dat)
```
- for multiple predictors/moderators:
    - main effects: `mods = ~ var1 + var2 + ...`
    - interactions: `mods = ~ var1 * var2 + ...`
  - character/factor variables:
    - are automatically dummy-coded
    - to remove the intercept: `mods = ~ var1 - 1`

```
> ### fit mixed-effects meta-regression model
> res <- rma(yi, vi, mods = ~ alloc + ablat, data=dat)
> res

Mixed-Effects Model (k = 13; tau^2 estimator: REML)

tau^2 (estimated amount of residual heterogeneity): 0.1446
tau (square root of estimated tau^2 value): 0.3883
I^2 (residual heterogeneity / unaccounted variability): 70.11%
H^2 (unaccounted variability / sampling variability): 3.35
R^2 (amount of heterogeneity accounted for): 53.84%

Test for Residual Heterogeneity:
QE(df = 9) = 26.2034, p-val = 0.0019

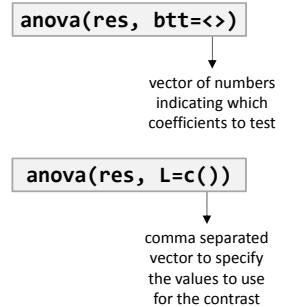
Test of Moderators (coefficient(s) 2,3,4):
QM(df = 3) = 11.0605, p-val = 0.0114

Model Results:

estimate se zval pval ci.lb ci.ub
intrcpt 0.2932 0.4050 0.7239 0.4691 -0.5006 1.0870
allocrandom -0.2675 0.3504 -0.7633 0.4453 -0.9543 0.4193
allocsystematic 0.0585 0.3795 0.1540 0.8776 -0.6854 0.8023
ablat -0.0273 0.0092 -2.9650 0.0030 -0.0453 -0.0092
```

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## Wald-Type Tests and Contrasts



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```
> ### omnibus test of the 'alloc' factor
> anova(res, btt=2:3)

Test of Moderators (coefficient(s) 2,3):
QM(df = 2) = 1.2850, p-val = 0.5260

> ### test random versus systematic allocation
> anova(res, L=c(0,1,-1,0))

Hypothesis:
1: allocrandom - allocsystematic = 0

Results:
estimate se zval pval
1: -0.3260 0.3104 -1.0501 0.2937
```

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```
> ### load data
> dat <- get(data(dat.konstantopoulos2011))
>
> ### show data
> dat
standardized mean differences and sampling variances
```

| district | school | study | year | yi   | vi          |
|----------|--------|-------|------|------|-------------|
| 1        | 11     | 1     | 1    | 1976 | -0.18 0.118 |
| 2        | 11     | 2     | 2    | 1976 | -0.22 0.118 |
| 3        | 11     | 3     | 3    | 1976 | 0.23 0.144  |
| 4        | 11     | 4     | 4    | 1976 | -0.30 0.144 |
| 5        | 12     | 1     | 5    | 1989 | 0.13 0.014  |
| 6        | 12     | 2     | 6    | 1989 | -0.26 0.014 |
| 7        | 12     | 3     | 7    | 1989 | 0.19 0.015  |
| 8        | 12     | 4     | 8    | 1989 | 0.32 0.024  |
| 9        | 18     | 1     | 9    | 1994 | 0.45 0.023  |
| 10       | 18     | 2     | 10   | 1994 | 0.38 0.043  |
| 11       | 18     | 3     | 11   | 1994 | 0.29 0.012  |
| 12       | ...    | ...   | ...  |      | ...         |
| 56       | 644    | 4     | 56   | 1994 | -0.05 0.067 |

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```
> ### fit standard random-effects model
> res <- rma(yi, vi, data = dat)
> res

Random-Effects Model (k = 56; tau^2 estimator: REML)

tau^2 (estimated amount of total heterogeneity): 0.0884
tau (square root of estimated tau^2 value): 0.2974
I^2 (total heterogeneity / total variability): 94.70%
H^2 (total variability / sampling variability): 18.89

Test for Heterogeneity:
Q(df = 55) = 578.8640, p-val < .0001

Model Results:

estimate se zval pval ci.lb ci.ub
0.1279 0.0439 2.9161 0.0035 0.0419 0.2139
```

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## The rma.mv() Function

- more flexible model fitting function, but must specify random effects manually
  - for now, replicate previous results
- ```
res <- rma.mv(yi, vi, random = ~ 1 | study,
                method = "REML", data = dat)
```
- `random = ~ 1 | study` adds a random effect for each level of the study variable
 - `method = "REML"` is default (other option: `ML`)

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```
> ### fit standard random-effects model with rma.mv()
> res <- rma.mv(yi, vi, random = ~ 1 | study, data = dat)
> res

Multivariate Meta-Analysis Model (k = 56; method: REML)

Variance Components:

          estim   sqrt  nlvls fixed factor
sigma^2    0.0884  0.2974     56    no  study

Test for Heterogeneity:
Q(df = 55) = 578.8640, p-val < .0001

Model Results:

estimate   se      zval    pval   ci.lb   ci.ub
0.1279  0.0439  2.9161  0.0035  0.0419  0.2139
```

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```
> ### load data
> dat <- get(data(dat.konstantopoulos2011))
>
> ### show data
> dat
```

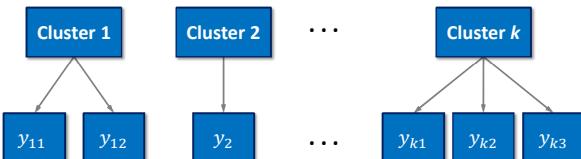
| | district | school | study | year | yi | vi |
|----|----------|--------|-------|------|-------|-------|
| 1 | 11 | 1 | 1 | 1976 | -0.18 | 0.118 |
| 2 | 11 | 2 | 2 | 1976 | -0.22 | 0.118 |
| 3 | 11 | 3 | 3 | 1976 | 0.23 | 0.144 |
| 4 | 11 | 4 | 4 | 1976 | -0.30 | 0.144 |
| 5 | 12 | 1 | 5 | 1989 | 0.13 | 0.014 |
| 6 | 12 | 2 | 6 | 1989 | -0.26 | 0.014 |
| 7 | 12 | 3 | 7 | 1989 | 0.19 | 0.015 |
| 8 | 12 | 4 | 8 | 1989 | 0.32 | 0.024 |
| 9 | 18 | 1 | 9 | 1994 | 0.45 | 0.023 |
| 10 | 18 | 2 | 10 | 1994 | 0.38 | 0.043 |
| 11 | 18 | 3 | 11 | 1994 | 0.29 | 0.012 |
| 12 | ... | ... | ... | ... | ... | ... |
| 56 | 644 | 4 | 56 | 1994 | -0.05 | 0.067 |

between 3 and
11 schools within
11 different
districts (56
studies in total)

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Multilevel Meta-Analytic Data

- multilevel structures can arise when we have multiple estimates for some higher clustering variable (paper, lab, research group, ...)



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Multilevel Random-Effects Model

$$\begin{aligned}
 y_{ij} = & \mu \quad \text{average true outcome} \\
 & + w_i \quad \text{random effect that makes the true outcomes for a particular cluster larger/smaller by some amount (heterogeneity between clusters)} \\
 & + u_{ij} \quad \text{random effect that makes one of the true outcomes within a particular cluster larger/smaller by some amount (heterogeneity within clusters)} \\
 & + e_{ij} \quad \text{sampling error}
 \end{aligned}$$

$$w_i \sim N(0, \sigma_w^2) \quad u_{ij} \sim N(0, \sigma_u^2) \quad e_{ij} \sim N(0, v_{ij})$$

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Marginal Variance-Covariance Matrix

$$Var \begin{bmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_2 \\ y_{31} \\ y_{32} \\ \vdots \end{bmatrix} = \begin{bmatrix} \sigma_B^2 + \sigma_w^2 + v_{11} & \sigma_B^2 & \sigma_B^2 \\ \sigma_B^2 + \sigma_w^2 + v_{12} & \sigma_B^2 & \\ \sigma_B^2 + \sigma_w^2 + v_{13} & & \\ \hline & \sigma_B^2 + \sigma_w^2 + v_2 & \\ & \sigma_B^2 + \sigma_w^2 + v_{31} & \sigma_B^2 \\ & \sigma_B^2 + \sigma_w^2 + v_{32} & \end{bmatrix} \dots$$

marginal variance-covariance matrix
with a block-diagonal structure

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The rma.mv() Function

- `rma.mv()` allows for the addition of multiple nested random effects
- `random = ~ 1 | var1/var2` adds a random effect for each level of `var1` and a random effect for each level of `var2` within each level of `var1`

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```

> ### multilevel random-effects model
> res <- rma.mv(yi, vi, data = dat,
+                 random = ~ 1 | district/school)
> res

Multivariate Meta-Analysis Model (k = 56; method: REML)

Variance Components:

          estim   sqrt  nvlvs  fixed      factor
sigma^2.1  0.0651  0.2551     11    no    district
sigma^2.2  0.0327  0.1809     56    no  district/school

Test for Heterogeneity:
Q(df = 55) = 578.8640, p-val < .0001

Model Results:

estimate   se   zval   pval   ci.lb   ci.ub
 0.1847  0.0846  2.1845  0.0289  0.0190  0.3504

```

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Correlation due to Multilevel Structure

- the multilevel structure implies that the true outcomes within a cluster are correlated:

$$\rho = \frac{\sigma_B^2}{\sigma_B^2 + \sigma_W^2}$$

- in example:

$$\hat{\rho} = \frac{0.0651}{0.0651 + 0.0327} = .67$$

- also note: $0.0651 + 0.0327 = 0.0978$

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Multivariate Parameterization

$$\begin{aligned}
y_{ij} &= \mu \quad \text{average true outcome} \\
&+ u_{ij} \quad \text{correlated random effects for the true outcomes within the same cluster} \\
&+ e_{ij} \quad \text{sampling error}
\end{aligned}$$

$$\begin{bmatrix} u_{i1} \\ u_{i2} \\ u_{i3} \end{bmatrix} \sim MVN \left(\begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \tau^2 & \rho\tau^2 & \rho\tau^2 \\ \rho\tau^2 & \tau^2 & \rho\tau^2 \\ \rho\tau^2 & \rho\tau^2 & \tau^2 \end{bmatrix} \right) \quad e_{ij} \sim N(0, v_{ij})$$

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Marginal Variance-Covariance Matrix

$$Var \begin{bmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{31} \\ y_{32} \\ \vdots \end{bmatrix} = \begin{bmatrix} \tau^2 + v_{11} & \rho\tau^2 & \rho\tau^2 & & & & \\ \rho\tau^2 & \tau^2 + v_{12} & \rho\tau^2 & & & & \\ \rho\tau^2 & \rho\tau^2 & \tau^2 + v_{13} & & & & \\ & & & \tau^2 + v_{21} & & & \\ & & & & \tau^2 + v_{31} & \rho\tau^2 & \\ & & & & \tau^2 + v_{32} & & \ddots \end{bmatrix}$$

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The rma.mv() Function

- random = ~ var1 | var2** adds correlated random effects for each level of **var1** within each level of **var2**
- note: **var1** must be a character/factor type variable (if it is not, use **factor()** function)

```

> ### fit multivariate random-effects model
> res <- rma.mv(yi, vi, random = ~ factor(school) | district,
+                 data = dat)
> res

Multivariate Meta-Analysis Model (k = 56; method: REML)

Variance Components:

outer factor: district      (nvlvs = 11)
inner factor: factor(school) (nvlvs = 11)

          estim   sqrt  fixed
tau^2     0.0978  0.3127    no
rho       0.6653        no

Test for Heterogeneity:
Q(df = 55) = 578.8640, p-val < .0001

Model Results:

estimate   se   zval   pval   ci.lb   ci.ub
 0.1847  0.0846  2.1845  0.0289  0.0190  0.3504

```

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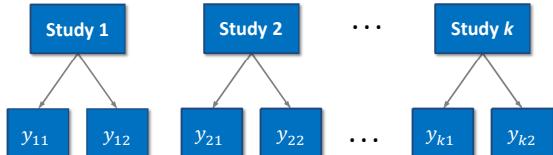
Notes

- models assume independent **sampling errors** within clusters (sensible if no overlap in the data/subjects used to compute outcomes)
- examples:
 - multiple independent studies reported in paper
 - multiple papers published by the same group
 - results reported for different subgroups
- but **true outcomes** within clusters may be more similar to each other than those from different clusters (correlated true outcomes)

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Multiple (Correlated) Outcomes

- multivariate data also arise when multiple outcomes are measured within the studies



note: not all studies have to measure all outcomes

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Multiple (Correlated) Outcomes

- since the outcomes are measured in the same subjects, the sampling errors are correlated
- true outcomes may also be correlated
- equations for the covariance between the sampling errors can be found in Gleser & Olkin (2009), Wei & Higgins (2013), Steiger (1980), ...

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Multivariate Random-Effects Model

$$\begin{aligned}
 y_{ij} = & \mu_j && \text{average true outcome for } j\text{th outcome} \\
 & + u_{ij} && \text{correlated random effects corresponding} \\
 & + e_{ij} && \text{correlated sampling errors of the observed} \\
 & && \text{outcomes for the same study (with known} \\
 & && \text{var-cov matrix)}
 \end{aligned}$$

$$\text{Var} \begin{bmatrix} u_{i1} \\ u_{i2} \end{bmatrix} = \begin{bmatrix} \tau_1^2 & \rho\tau_1\tau_2 \\ \rho\tau_1\tau_2 & \tau_2^2 \end{bmatrix} \quad \text{Var} \begin{bmatrix} e_{i1} \\ e_{i2} \end{bmatrix} = \begin{bmatrix} v_{i1} & \text{cov}_i \\ & v_{i2} \end{bmatrix}$$

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```

> ### load data
> dat <- get(data(dat.berkey1998))
>
> ### show data
> dat
      trial   author year ni outcome   yi   v1i   v2i
1       1 Pihlstrom et al. 1983 14   PD 0.47 0.0075 0.0030
2       1 Pihlstrom et al. 1983 14   AL -0.32 0.0030 0.0077
3       2 Lindhe et al. 1982 15   PD 0.20 0.0057 0.0009
4       2 Lindhe et al. 1982 15   AL -0.60 0.0009 0.0008
5       3 Knowles et al. 1979 78   PD 0.40 0.0021 0.0007
6       3 Knowles et al. 1979 78   AL -0.12 0.0007 0.0014
7       4 Ramfjord et al. 1987 89   PD 0.26 0.0029 0.0009
8       4 Ramfjord et al. 1987 89   AL -0.31 0.0009 0.0015
9       5 Becker et al. 1988 16   PD 0.56 0.0148 0.0072
10      5 Becker et al. 1988 16   AL -0.39 0.0072 0.0304
  
```

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```

> ### construct var-cov matrix of the sampling errors
> V <- split(dat[,c("v1i","v2i")], dat$trial)
> V <- lapply(V, as.matrix)
> V <- bldiag(V)
> V
      [,1]   [,2]   [,3]   [,4]   [,5]   [,6]   [,7]   [,8]   [,9]   [,10]
[1,] 0.0075 0.0030 0.0000 0.0000 ... ... ... ... ... ...
[2,] 0.0030 0.0077 0.0000 0.0000 ... ... ... ... ... ...
[3,] 0.0000 0.0000 0.0057 0.0009 ... ... ... ... ... ...
[4,] 0.0000 0.0000 0.0009 0.0008 ... ... ... ... ... ...
[5,] ... ... ... ... ... ... ... ... ... ...
[6,] ... ... ... ... ... ... ... ... ...
[7,] ... ... ... ... ... ... ...
[8,] ... ... ... ... ...
[9,] ... ... ... ...
[10,] ... ... ... ...

```

| | |
|--------|--------|
| 0.0148 | 0.0072 |
| 0.0072 | 0.0304 |

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The rma.mv() Function

```
name of object with      name of factor to
the var-cov matrix of    indicate the outcome
the sampling errors      (and remove intercept)
↑                         ↑
res <- rma.mv(yi, V, mods = ~ outcome - 1,
               random = ~ outcome | study,
               struct = "UN", data = dat)
↓
structure of var-cov matrix of the
random effects (UN = unstructured)
```

```
> ### multivariate random-effects model
> res <- rma.mv(yi, V, mods = ~ outcome - 1, data = dat,
+                 random = ~ outcome | trial, struct = "UN")
> res
Multivariate Meta-Analysis Model (k = 10; method: REML)

Variance Components:

outer factor: trial (nlevels = 5)
inner factor: outcome (nlevels = 2)

estim   sqrt k.lvl fixed level
tau^2.1 0.0327 0.1807 5 no AL
tau^2.2 0.0117 0.1083 5 no PD
rho     0.6088

Test for Residual Heterogeneity:
QE(df = 8) = 128.2267, p-val < .0001

Model Results:

estimate   se   zval   pval ci.lb ci.ub
outcomeAL -0.3392 0.0879 -3.8588 0.0001 -0.5115 -0.1669
outcomePD  0.3534 0.0588  6.0056 <.0001  0.2381  0.4688
```

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Random Effects Structures

struct="CS"
(this is the default)

$$\begin{bmatrix} \tau^2 & \rho\tau^2 & \rho\tau^2 & \rho\tau^2 \\ \tau^2 & \tau^2 & \rho\tau^2 & \rho\tau^2 \\ \tau^2 & \rho\tau^2 & \tau^2 & \rho\tau^2 \\ \tau^2 & \rho\tau^2 & \tau^2 \end{bmatrix}$$

struct="HCS"

$$\begin{bmatrix} \tau_1^2 & \rho\tau_1\tau_2 & \rho\tau_1\tau_3 & \rho\tau_1\tau_4 \\ \tau_2^2 & \tau_2^2 & \rho\tau_2\tau_3 & \rho\tau_2\tau_4 \\ \tau_3^2 & \rho\tau_3\tau_4 & \tau_3^2 & \tau_3^2 \\ \tau_4^2 & \tau_4^2 & \tau_4^2 & \tau_4^2 \end{bmatrix}$$

struct="UN"

$$\begin{bmatrix} \tau_1^2 & \rho_{12}\tau_1\tau_2 & \rho_{13}\tau_1\tau_3 & \rho_{14}\tau_1\tau_4 \\ \tau_2^2 & \rho_{23}\tau_2\tau_3 & \rho_{24}\tau_2\tau_4 & \tau_2^2 \\ \tau_3^2 & \rho_{34}\tau_3\tau_4 & \tau_3^2 & \tau_3^2 \\ \tau_4^2 & \tau_4^2 & \tau_4^2 & \tau_4^2 \end{bmatrix}$$

(for two outcomes, "UN" and "HCS" are the same)

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```
> ### contrast for differences in outcomes
> anova(res, L=c(1,-1))
```

Hypothesis:

1: outcomeAL - outcomePD = 0

Results:

| | estimate | se | zval | pval |
|----|----------|--------|---------|--------|
| 1: | -0.6926 | 0.0744 | -9.3120 | <.0001 |

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Multiple Time Points

- multivariate data also arise when an outcome is measured at multiple time points
- the sampling errors will again be correlated
- true outcomes may also be correlated
- can consider auto-regressive structures for the sampling errors and random effects (Ishak et al., 2007; Trikalinos & Olkin, 2012)

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Random Effects Structures

struct="AR"

$$\begin{bmatrix} \tau^2 & \rho\tau^2 & \rho^2\tau^2 & \rho^3\tau^2 \\ \tau^2 & \tau^2 & \rho\tau^2 & \rho^2\tau^2 \\ \tau^2 & \rho\tau^2 & \tau^2 & \rho\tau^2 \\ \tau^2 & \rho\tau^2 & \tau^2 & \tau^2 \end{bmatrix}$$

struct="HAR"

$$\begin{bmatrix} \tau_1^2 & \rho\tau_1\tau_2 & \rho^2\tau_1\tau_3 & \rho^3\tau_1\tau_4 \\ \tau_2^2 & \tau_2^2 & \rho\tau_2\tau_3 & \rho^2\tau_2\tau_4 \\ \tau_3^2 & \tau_3^2 & \rho\tau_3\tau_4 & \tau_3^2 \\ \tau_4^2 & \tau_4^2 & \tau_4^2 & \tau_4^2 \end{bmatrix}$$

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```

> ### load data
> dat <- get(data(dat.ishak2007))
>
> ### create long format dataset
> dat.long <- reshape(dat, direction="long", idvar="study",
+   v.names=c("yi","vi"),
+   varying=list(c(2,4,6,8), c(3,5,7,9)))
> dat.long <- dat.long[order(dat.long$study, dat.long$time),]
> rownames(dat.long) <- 1:nrow(dat.long)
>
> ### remove missing measurement occasions from dat.long
> is.miss <- is.na(dat.long$yi)
> dat.long <- dat.long[!is.miss,]
>
> ### construct full (block diagonal) V matrix with AR(1) structure
> rho.within <- .97
> V <- lapply(split(with(dat, cbind(vii, v2i, v3i, v4i)),
+   dat$study), diag)
> V <- lapply(V, function(v) sqrt(v) *%
+   toeplitz(ARMAacf(ar=rho.within, lag.max=3)) %*% sqrt(v))
> V <- bldiag(V)
> V <- V[!is.miss,!is.miss]

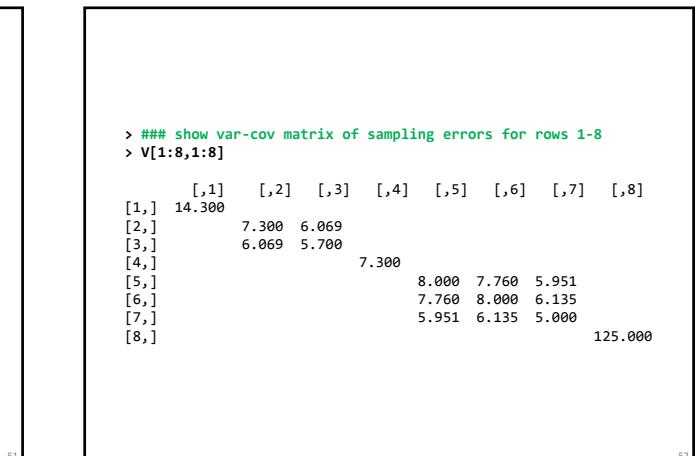
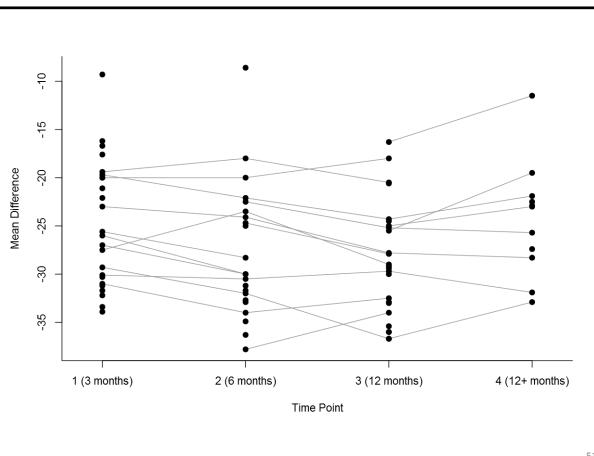
```

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> ### show data
> dat.long

| study | time | yi | vi | |
|-------|--------------------|-----|-------|-------|
| 1 | Alegret (2001) | 1 | -33.4 | 14.3 |
| 5 | Barichella (2003) | 1 | -20.0 | 7.3 |
| 7 | Barichella (2003) | 3 | -30.0 | 5.7 |
| 9 | Berney (2002) | 1 | -21.1 | 7.3 |
| 13 | Burchiel (1999) | 1 | -20.0 | 8.0 |
| 14 | Burchiel (1999) | 2 | -20.0 | 8.0 |
| 15 | Burchiel (1999) | 3 | -18.0 | 5.0 |
| 18 | Chen (2003) | 2 | -32.9 | 125.0 |
| 91 | ... | ... | ... | ... |
| 173 | Vingerhoets (2002) | 1 | -19.7 | 18.5 |
| 174 | Vingerhoets (2002) | 2 | -22.1 | 18.1 |
| 175 | Vingerhoets (2002) | 3 | -24.3 | 18.2 |
| 176 | Vingerhoets (2002) | 4 | -21.9 | 16.7 |
| 178 | Volkman (2001) | 2 | -37.8 | 20.9 |
| 179 | Volkman (2001) | 3 | -34.0 | 26.4 |
| 181 | Weselburger (2002) | 1 | -22.1 | 40.8 |

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```

> ### multivariate model with HAR(1) structure
> res <- rma.mv(yi, V, mods = ~ factor(time) - 1, data = dat.long,
+   random = ~ time | study, struct = "HAR")
> res

Multivariate Meta-Analysis Model (k = 82; method: REML)

Variance Components:

estim    sqrt  k.lvl fixed  level
tau^2.1  22.724  4.767    24 no     1
tau^2.2  33.755  5.810    22 no     2
tau^2.3  26.167  5.115    25 no     3
tau^2.4  31.207  5.586    11 no     4
rho      0.883          no

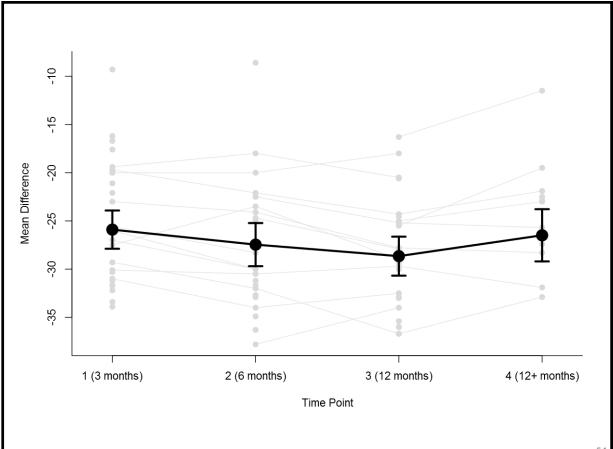
Test for Residual Heterogeneity:
QE(df = 78) = 856.164, p-val < .001

Model Results:

estimate   se    zval  pval ci.lb ci.ub
factor(time)1 -25.905 1.012 -25.605 <.001 -27.888 -23.922
factor(time)2 -27.461 1.141 -24.072 <.001 -29.697 -25.225
factor(time)3 -28.656 1.032 -27.756 <.001 -30.680 -26.633
factor(time)4 -26.494 1.382 -19.172 <.001 -29.202 -23.785

```

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```
> ### comparison of effects between pairs of time points
> anova(res, L=rbind(c(1,-1,0,0), c(1,0,-1,0), c(1,0,0,-1),
+ c(0,1,-1,0), c(0,1,0,-1),
+ c(0,0,1,-1)))
```

Hypotheses:

- 1: factor(time)1 - factor(time)2 = 0
- 2: factor(time)1 - factor(time)3 = 0
- 3: factor(time)1 - factor(time)4 = 0
- 4: factor(time)2 - factor(time)3 = 0
- 5: factor(time)2 - factor(time)4 = 0
- 6: factor(time)3 - factor(time)4 = 0

Results:

| | estimate | se | zval | pval | |
|----|----------|-------|--------|-------|--------|
| 1: | 1.556 | 0.755 | 2.061 | 0.039 | 1 vs 2 |
| 2: | 2.751 | 0.859 | 3.204 | 0.001 | 1 vs 3 |
| 3: | 0.589 | 1.273 | 0.462 | 0.644 | 1 vs 4 |
| 4: | 1.195 | 0.761 | 1.569 | 0.117 | 2 vs 3 |
| 5: | -0.967 | 1.217 | -0.795 | 0.427 | 2 vs 4 |
| 6: | -2.163 | 0.968 | -2.235 | 0.025 | 3 vs 4 |

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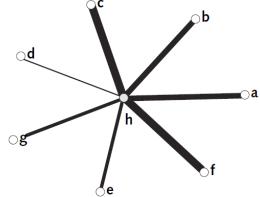
Network Meta-Analysis

- often there are multiple treatments available for the same condition/disease
- studies comparing the effectiveness of these treatments form a network of comparisons
- some of the goals:
 - synthesize evidence provided by all studies and comparisons in one parsimonious model
 - obtain indirect evidence about comparisons that have not been examined head-to-head
 - determine a hierarchy of treatment effectiveness

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Star-Shaped Networks

Second-generation antiepileptic drugs in partial epilepsy

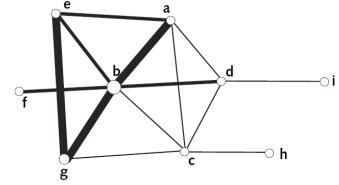


a: levetiracetam, b: gabapentin, c: lamotrigine, d: oxcarbazepine, e: tiagabine, f: topiramate, g: zonisamide, h: placebo

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Complex Networks

Chemotherapy regimens for ovarian cancer



a: platinum monotherapy, b: platinum-based combination, c: taxane monotherapy, d: platinum + taxane-based combination, e: nonplatinum/nontaxane monotherapy, f: platinum-based combination (ip), g: nonplatinum/nontaxane combination, h: taxane-based combination, i: platinum/taxane-based combination (ip)

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Network Meta-Analysis

- can analyze such data with appropriate multilevel/multivariate models
- two general approaches: arm- vs. contrast-based model (e.g., Salanti et al., 2008)
- errors are correlated in contrast-based model for studies with more than two groups
- equations for the correlation between the sampling errors can be found in Gleser and Olkin (2009) and several other papers

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Arm-Based Network Meta-Analysis

$$y_{ij} = \beta_0 + \beta_1 T_{i1} + \dots + \beta_p T_{ip} \quad (T_{ij} = \text{treatment indicators})$$

+ w_i random effect that makes the true outcomes for a particular study larger/smaller by some amount (between-study heterogeneity)

+ u_{ij} random effect that makes one of the true outcomes within a particular study larger/smaller by some amount (between-treatment heterogeneity)

+ e_{ij} sampling error

$$w_i \sim N(0, \sigma_w^2) \quad u_{ij} \sim N(0, \sigma_u^2) \quad e_{ij} \sim N(0, v_{ij})$$

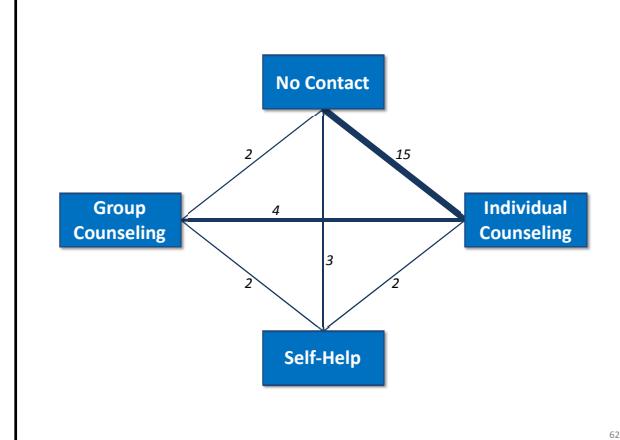
60

```

> ### load data
> dat <- get(data(dat.hasselblad1998))
>
> ### calculate log odds for each study arm
> dat <- escalc(measure="PLO", xi=xi, ni=ni, data=dat)
>
> ### show data
> dat
log odds and corresponding
sampling variances

```

| | id | study | trt | xi | ni | yi | vi |
|----|-----|-------|----------------|-----|------|--------|-------|
| 1 | 1 | 1 | no_contact | 75 | 731 | -2.169 | 0.015 |
| 2 | 2 | 1 | ind_counseling | 363 | 714 | 0.034 | 0.006 |
| 3 | 3 | 2 | no_contact | 9 | 140 | -2.678 | 0.119 |
| 4 | 4 | 2 | ind_counseling | 23 | 140 | -1.627 | 0.052 |
| 5 | 5 | 2 | grp_counseling | 10 | 138 | -2.549 | 0.108 |
| 6 | 6 | 3 | no_contact | 2 | 106 | -3.951 | 0.510 |
| 7 | 7 | 3 | ind_counseling | 9 | 205 | -3.081 | 0.116 |
| 8 | ... | ... | ... | ... | ... | ... | ... |
| 9 | 49 | 24 | no_contact | 69 | 1177 | -2.776 | 0.015 |
| 10 | 50 | 24 | ind_counseling | 54 | 888 | -2.737 | 0.020 |



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```

> ### convert trt variable to factor with desired ordering of levels
> dat$trt <- factor(dat$trt, levels=c("no_contact", "self_help",
+ "ind_counseling", "grp_counseling"))
>
> ### network meta-analysis using a multilevel model
> res <- rma.mv(yi, vi, mods = ~ trt, data = dat,
+                 random = ~ 1 | study/trt)
> res

Multivariate Meta-Analysis Model (k = 50; method: REML)

Variance Components:

estim   sqrt  nlvs  fixed     factor
sigma^2.1 0.195 0.441    24    no  study
sigma^2.2 0.249 0.499    50    no  study/trt

Test of Moderators (coefficient(s) 2,3,4):
QM(df = 3) = 19.441, p-val < .001

Model Results:

estimate      se      zval    pval    ci.lb    ci.ub
intrcpt     -2.456 0.174  -14.129 <.001  -2.796  -2.115
trt self_help  0.501 0.302   1.657 0.098  -0.092  1.093
trt ind_counseling  0.777 0.196   3.969 <.001  0.393  1.161
trt grp_counseling  1.056 0.324   3.259 0.001  0.421  1.691

```

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```

> ### pairwise odds ratios of interventions versus no contact
> predict(res, newmods=diag(3),
+          intercept=FALSE, transf=exp, digits=2)

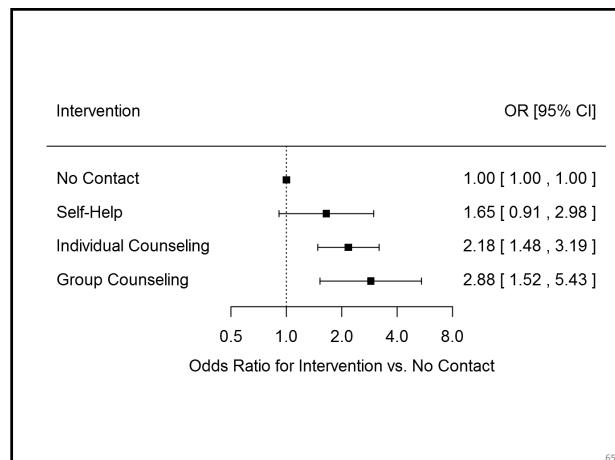
pred ci.lb ci.ub cr.lb cr.ub
1 1.65 0.91 2.98 0.39 6.92  Self-Help versus No Contact
2 2.18 1.48 3.19 0.56 8.49  Individual Counseling versus No Contact
3 2.88 1.52 5.43 0.67 12.29  Group Counseling versus No Contact

> ### pairwise odds ratios comparing interventions
> predict(res, newmods=rbind(c(-1,1,0), c(-1,0,1), c(0,-1,1)),
+          intercept=FALSE, transf=exp, digits=2)

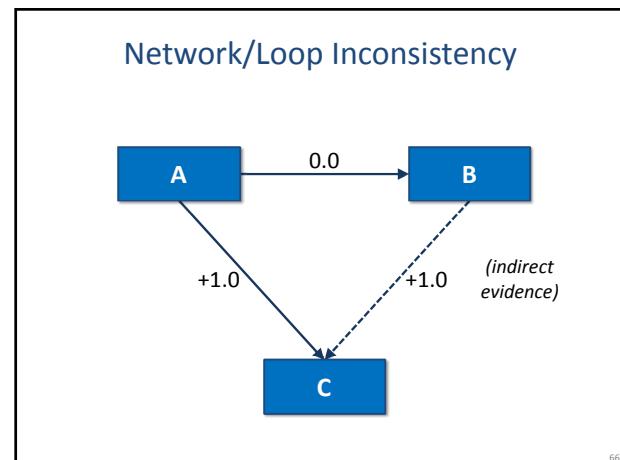
pred ci.lb ci.ub cr.lb cr.ub
1 1.32 0.73 2.39 0.31 5.54  Individual Counseling versus Self-Help
2 1.74 0.84 3.62 0.39 7.79  Group Counseling versus Self-Help
3 1.32 0.72 2.43 0.31 5.58  Group versus Individual Counseling

```

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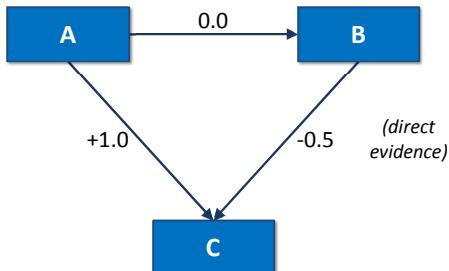


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Network/Loop Inconsistency



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Dealing with Inconsistency

- restrict analysis to a subset of studies providing consistent evidence
- try to account for it based moderators
- model it (various proposals)

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Final Notes

- `rma.mv()` allows for an arbitrary number of random effects of the form `~ 1 | factor` (allows for 3/4/...-level models and crossed random effects)
- up to 2 terms of the form `~ inner | outer`
- can also specify a known correlation matrix corresponding to a `~ 1 | factor` term (e.g., for phylogenetic meta-analyses)
- website: <http://www.metafor-project.org>

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- Adams, D. C. (2008). Phylogenetic meta-analysis. *Evolution*, 62(3), 567-572.
- Berkey, C. S., Hoaglin, D. C., Antczak-Bouckoms, A., Mosteller, F., & Colditz, G. A. (1998). Meta-analysis of multiple outcomes by regression with random effects. *Statistics in Medicine*, 17(22), 2537-2550.
- Colditz, G. A., Brewer, T. F., Berkey, C. S., Wilson, M. E., Burdick, E., Fineberg, H. V., & Mosteller, F. (1994). Efficacy of BCG vaccine in the prevention of tuberculosis: Meta-analysis of the published literature. *Journal of the American Medical Association*, 271(9), 698-702.
- Cooper, H., Valentine, J. C., Charlton, K., & Nelson, A. (2005). The effects of modified school calendars on student achievement and on school and community attitudes. *Review of Educational Research*, 75(1), 1-52.
- Gleser, L. J., & Olkin, I. (2009). Stochastically dependent effect sizes. In H. Cooper, L. V. Hedges, & J. C. Valentine (Eds.), *The handbook of research synthesis and meta-analysis* (2nd ed., pp. 357-376). New York: Russell Sage Foundation.
- Ishak, K. J., Platt, R. W., Joseph, L., Hanley, J. A., & Caro, J. J. (2007). Meta-analysis of longitudinal studies. *Clinical Trials*, 4(5), 525-539.
- Konstantopoulos, S. (2011). Fixed effects and variance components estimation in three-level meta-analysis. *Research Synthesis Methods*, 2(1), 61-76.
- Lajeunesse, M. J. (2009). Meta-analysis and the comparative phylogenetic method. *The American Naturalist*, 174(3), 369-381.
- Nakagawa, S., & Santos, E. S. A. (2012). Methodological issues and advances in biological meta-analysis. *Evolutionary Ecology*, 26(5), 1253-1274.
- Salanti, G., Higgins, J. P. T., Ades, A. E., & Ioannidis, J. P. A. (2008). Evaluation of networks of randomized trials. *Statistical Methods in Medical Research*, 17(3), 279-301.
- Steiger, J. H. (1980). Tests for comparing elements of a correlation matrix. *Psychological Bulletin*, 87(2), 245-251.
- Trikalinos, T. A., & Olkin, I. (2012). Meta-analysis of effect sizes reported at multiple time points: A multivariate approach. *Clinical Trials*, 9(5), 610-620.
- Viechtbauer, W. (2010). Conducting meta-analyses in R with the metafor package. *Journal of Statistical Software*, 36(3), 1-48.
- Wei, Y., & Higgins, J. P. (2013). Estimating within-study covariances in multivariate meta-analysis with multiple outcomes. *Statistics in Medicine*, 32(7), 1191-1205.

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Thank You!

Questions? Comments?

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